

# Lectures on Stochastic Stability

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## Lecture 3

### Fluid Approximation Approach and Induced Vector Fields

We consider two stability methods for Markov chains based on the drift analysis.

#### 1 Fluid approximation approach

In this section, we give essentially an application of Lyapunov methods to the so-called stability via fluid limits, a technique which became popular in the 90's. Roughly speaking, fluid approximation refers to a functional law of large numbers which may be formulated for large classes of Markovian and non-Markovian systems. Instead of trying to formulate the technique very generally, we focus on a quite important class of stochastic models, namely, multi-class networks. For statements and proofs of the functional approximation theorems used here, the reader may consult the texts of Chen and Yao [4], Whitt [11] and references therein.

##### 1.1 Exemplifying the technique in a simple case

To exemplify the technique we start with a GI/GI/1 queue with general non-idling, work-conserving, non-preemptive service discipline.<sup>1</sup> Let  $Q(t)$ ,  $\chi(t)$ ,  $\psi(t)$  be, respectively, the number of customers in the system, remaining service time of customer at the server (if any), and remaining interarrival time, at time  $t$ . The three quantities, together, form a Markov process. We will scale the whole process by

$$N = Q(0) + \chi(0) + \psi(0).$$

Although it is tempting, based on a functional law of large numbers (FLLN), to assert that  $Q(Nt)/N$  has a limit, as  $N \rightarrow \infty$ , this is not quite right, unless we specify how the individual constituents of  $N$  behave. So, we assume that<sup>2</sup>

$$Q(0) \sim c_1 N, \quad \chi(0) \sim c_2 N, \quad \psi(0) \sim c_3 N, \quad \text{as } N \rightarrow \infty,$$

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<sup>1</sup>This means that when a customer arrives at the server with  $\sigma$  units of work, then the server works with the customer without interruption, and it takes precisely  $\sigma$  time units for the customer to leave.

<sup>2</sup>Hence, strictly speaking, we should denote the process by an extra index  $N$  to denote this dependence, i.e., write  $Q^{(N)}(t)$  in lieu of  $Q(t)$ , but, to save space, we shall not do so.

where  $c_1 + c_2 + c_3 = 1$ . Then

$$\frac{Q(Nt)}{N} \rightarrow \bar{Q}(t), \quad \text{as } N \rightarrow \infty,$$

uniformly on compact<sup>3</sup> sets of  $t$ , a.s., i.e.,

$$\lim_{N \rightarrow \infty} \mathbf{P} \left( \sup_{0 \leq t \leq T} |Q(kt)/k - \bar{Q}(t)| > \varepsilon, \text{ for some } k > N \right) = 0, \quad \text{for all } T, \varepsilon > 0.$$

The function  $\bar{Q}$  is defined by:

$$\bar{Q}(t) = \begin{cases} c_1, & t < c_3 \\ c_1 + \lambda(t - c_3), & c_3 \leq t < c_2, \quad \text{if } c_3 \leq c_2, \\ (c_1 + \lambda(c_2 - c_3) + (\lambda - \mu)(t - c_2))^+, & t \geq c_2 \end{cases}$$

$$\bar{Q}(t) = \begin{cases} c_1, & t < c_2 \\ c_1 - \mu(t - c_2), & c_2 \leq t < c_3, \quad \text{if } c_2 < c_3. \\ ((c_1 - \mu(c_3 - c_2))^+ + (\lambda - \mu)(t - c_3))^+, & t \geq c_3 \end{cases}$$

It is clear that  $\bar{Q}(t)$  is the difference between two continuous, piecewise linear, and increasing functions. We shall not prove this statement here, because it is more than what we need: indeed, as will be seen later, the full functional law of large numbers tells a more detailed story; all we need is the fact that there is a  $t_0 > 0$  that does not depend on the  $c_i$ , so that  $\bar{Q}(t) = 0$  for all  $t > t_0$ , provided we assume that  $\lambda < \mu$ . This can be checked directly from the formula for  $\bar{Q}$ . (On the other hand, if  $\lambda > \mu$ , then  $\bar{Q}(t) \rightarrow \infty$ , as  $t \rightarrow \infty$ .)

To translate this FLLN into a Lyapunov function criterion, we use an embedding technique: we sample the process at the  $n$ -th arrival epoch  $T_n$ . We take for simplicity  $T_0 = 0$ . It is clear that then we here can omit the state component  $\psi$ , because

$$X_n := (Q_n, \chi_n) := (Q(T_n), \chi(T_n))$$

is a Markov chain with state space  $\mathcal{X} = \mathbb{Z}_+ \times \mathbb{R}_+$ . So, we assume  $N = Q_0 + \chi_0 \rightarrow \infty$  and

$$Q(0) \sim c_1 N, \quad \chi(0) \sim c_2 N, \quad \text{with } c_1 + c_2 = 1.$$

Using another FLLN for the random walk  $T_n$ , namely,

$$\frac{T_{[N\lambda t]}}{N} \rightarrow t, \quad \text{as } N \rightarrow \infty, \quad \text{u.o.c., a.s.,}$$

we obtain, using the usual method via the continuity of the composition mapping,

$$\frac{Q_{[N\lambda t]}}{N} \rightarrow (c_1 + \lambda \min(t, c_2) + (\lambda - \mu)(t - c_2)^+)^+, \quad \text{as } N \rightarrow \infty, \quad \text{u.o.c., a.s..}$$

Under the stability condition  $\lambda < \mu$  and a uniform integrability (which shall be proved below) we have:

$$\frac{\mathbf{E}Q_{[N\lambda t]}}{N} \rightarrow 0, \quad \frac{\mathbf{E}\chi_{[N\lambda t]}}{N} \rightarrow 0, \quad \text{as } N \rightarrow \infty, \quad \text{for } t \geq t_0.$$

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<sup>3</sup>We abbreviate this as ‘‘u.o.c.’’; it is the convergence also know as compact convergence. See the Appendix for some useful properties of the u.o.c. convergence.

In particular there is  $N_0$ , so that  $\mathbf{E}Q_{[2N\lambda t_0]} + \mathbf{E}\chi_{[2N\lambda t_0]} \leq N/2$  for all  $N > N_0$ . Also, the same uniform integrability condition, allows us to find a constant  $C$  such that  $\mathbf{E}Q_{[2N\lambda t_0]} + \mathbf{E}\chi_{[2N\lambda t_0]} \leq C$  for all  $N \leq N_0$ . To translate this into the language of a Lyapunov criterion, let  $x = (q, \chi)$  denote a generic element of  $\mathcal{X}$ , and consider the functions

$$V(q, \chi) = q + \chi, \quad g(q, \chi) = 2q\lambda t_0, \quad h(q, \chi) = (1/2)q - C\mathbf{1}(q \leq N_0).$$

The last two inequalities can then be written as  $\mathbf{E}_x(V(X_{g(x)}) - V(X_0)) \leq -h(x)$ ,  $x \in \mathcal{X}$ . It is easy to see that the function  $V, g, h$  satisfy conditions (L0)-(L4) from the previous lecture. Thus the main Theorem 2 of the previous lecture shows that the set  $\{x \in \mathcal{X} : V(x) = q + \chi \leq N_0\}$  is positive recurrent.

## 1.2 Fluid limit stability criterion for multiclass queueing networks

We now pass on to multiclass queueing networks. Rybko and Stolyar [10] first applied the method to a two-station, two-class network. Dai [5] generalised the method and his paper established and popularised it. Meanwhile, it became clear that the natural stability conditions<sup>4</sup> may not be sufficient for stability and several examples were devised to exemplify this phenomena; see, e.g., again the paper by Rybko and Stolyar or the paper by Bramson [2] which gives an example of a multiclass network which is unstable under the natural stability conditions (the local traffic intensity at each node is below 1), albeit operating under the “simplest” possible discipline (FIFO).

To describe a multiclass queueing network, we let  $\{1, \dots, K\}$  be a set of customer classes and  $\{1, \dots, J\}$  a set of stations. Each station  $j$  is a single-server service facility that serves customers from the set of classes  $c(j)$  according to a non-idling, work-conserving, non-preemptive, but otherwise general, service discipline. It is assumed that  $c(j) \cap c(i) = \emptyset$  if  $i \neq j$ . There is a single arrival stream<sup>5</sup>, denoted by  $A(t)$ , which is the counting process of a renewal process, viz.,

$$A(t) = \mathbf{1}(\psi(0) \leq t) + \sum_{n \geq 1} \mathbf{1}(\psi(0) + T_n \leq t),$$

where  $T_n = \xi_1 + \dots + \xi_n$ ,  $n \in \mathbb{N}$ , and the  $\{\xi_n\}$  are i.i.d. positive r.v.’s with  $E\xi_1 = \lambda^{-1} \in (0, \infty)$ . The interpretation is that  $\psi(0)$  is the time required for customer 1 to enter the system, while  $T_n$  is the arrival time of customer  $n \in \mathbb{N}$ . (Artificially, we may assume that there is a customer 0 at time 0.) To each customer class  $k$  there corresponds a random variable  $\sigma_k$  used as follows: when a customers from class  $k$  is served, then its service time is an independent copy of  $\sigma_k$ . We let  $\mu_k^{-1} = \mathbf{E}\sigma_k$ . Routing at the arrival point is done according to probabilities  $p_k$ , so that an arriving customer becomes of class  $k$  with probability  $p_k$ . Routing in the network is done so that a customer finishing service from class  $k$  joins class  $\ell$  with probability  $p_{k,\ell}$ , and leaves the network with probability  $p_{k,\infty} = 1 - \sum_{\ell} p_{k,\ell}$ .

**Examples.** 1. Jackson-type (or generalised Jackson) network: there is one-to-one correspondence between stations and customer classes.

<sup>4</sup>By the term “natural stability conditions” in a work-conserving, non-idling queueing network we refer to the condition that says that the rate at which work is brought into a node is less than the processing rate.

<sup>5</sup>But do note that several authors consider many independent arrival streams

2. Kelly network. There are several deterministic routes, say,  $(j_{1,1}, \dots, j_{i,r_1}), \dots, (j_{m,1}, \dots, j_{m,r_m})$  where  $j_{i,r}$  are stations numbers. Introduce  $K = \sum_{q=1}^m r_q$  customers classes numbered  $1, \dots, K$  and let

$$p_{k,k+1} = 1 \quad \text{for } k \neq r_1, r_1 + r_2, \dots$$

and

$$p_{k,\infty} = 1 \quad \text{for } k = r_1, r_1 + r_2, \dots$$

Return to the general framework. Let  $A_k(t)$  be the cumulative arrival process of class  $k$  customers from the outside world. Let  $D_k(t)$  be the cumulative departure process from class  $k$ . The process  $D_k(t)$  counts the total number of departures from class  $k$ , both those that are recycled within the network and those who leave it. Of course, it is the specific service policies that will determine  $D_k(t)$  for all  $k$ . If we introduce i.i.d. routing variables  $\{\alpha_k(n), n \in \mathbb{N}\}$  so that  $\mathbf{P}(\alpha_k(n) = \ell) = p_{k\ell}$ , then we may write the class- $k$  dynamics as:

$$Q_k(t) = Q_k(0) + A_k(t) + \sum_{\ell=1}^K \sum_{n=1}^{D_\ell(t)} \mathbf{1}(\alpha_\ell(n) = k) - D_k(t).$$

In addition, a number of other equations are satisfied by the system: Let  $W^j(t)$  be the workload in station  $j$ . Let  $C_{jk} = \mathbf{1}(k \in c(j))$ . And let  $V(n) = \sum_{m=1}^n \sigma_k(n)$  be the sum of the service times brought by the first  $n$  class- $k$  customers. Then the total work brought by those customers up to time  $t$  is  $V_k(Q_k(0) + A_k(t))$ , and part of it, namely  $\sum_k C_{jk} V_k(Q_k(0) + A_k(t))$  is gone to station  $j$ . Hence the work present in station  $j$  at time  $t$  is

$$W^j(t) = \sum_k C_{jk} V_k(Q_k(0) + A_k(t)) - t + Y^j(t),$$

where  $Y^j(t)$  is the idleness process, viz.,

$$\int W^j(t) dY^j(t) = 0.$$

The totality of the equations above can be thought of as having inputs (or “primitives”) the  $\{A_k(t)\}$ ,  $\{\sigma_k(n)\}$  and  $\{\alpha_k(n)\}$ , and are to be “solved” for  $\{Q_k(t)\}$  and  $\{W^j(t)\}$ . However, they are not enough: more equations are needed to describe how the server spends his service effort to various customers, i.e, we need policy-specific equations; see, e.g., [4].

Let  $Q^j(t) = \sum_{k \in c(j)} Q_k(t)$ . Let  $\zeta_m^j(t)$  be the class of the  $m$ -th customer in the queue of station  $j$  at time  $t$ , so that  $\zeta^j(t) := (\zeta_1^j(t), \zeta_2^j(t), \dots, \zeta_{Q^j(t)}^j(t))$  is an array detailing the classes of all the  $Q^j(t)$  customers present in the queue of station  $j$  at time  $t$ , where the leftmost one refers to the customer receiving service (if any) and the rest to the customers that are waiting in line. Let also  $\chi^j(t)$  be the remaining service time of the customer receiving service. We refer to  $X^j(t) = (Q^j(t), \zeta^j(t), \chi^j(t))$  as the state<sup>6</sup> of station  $j$ . Finally, let  $\psi(t)$  be such that  $t + \psi(t)$  is the time of the first exogenous customer arrival after  $t$ . Then the most detailed information that will result in a Markov process in continuous time is  $X(t) := (X^1(t), \dots, X^J(t); \psi(t))$ . To be pedantic, we note that the state space of  $X(t)$  is  $\mathcal{X} = (\mathbb{Z}_+ \times K^* \times \mathbb{R}_+)^J \times \mathbb{R}_+$ , where  $K^* = \cup_{n=0}^{\infty} \{1, \dots, K\}^n$ , with  $\{1, \dots, K\}^0 = \{\emptyset\}$ , i.e.,  $\mathcal{X}$  is a horribly looking creature—a Polish space nevertheless.

<sup>6</sup>Note that the first component is, strictly speaking, redundant as it can be read from the length of the array  $\zeta^j(t)$ .

We now let

$$N = \sum_{j=1}^J (Q^j(0) + \chi^j(0)) + \psi(0),$$

and consider the system parametrised by this parameter  $N$ . While it is clear that  $A(Nt)/N$  has a limit as  $N \rightarrow \infty$ , it is not clear at all that so do  $D_k(Nt)/N$ . The latter depends on the service policies, and, even if a limit exists, it may exist only along a certain subsequence. This was seen even in the very simple case of a single server queue.

To precise about the notion of limit point used in the following definition, we say that  $\bar{X}(\cdot)$  is a limit point of  $X_N(\cdot)$  if there exists a deterministic subsequence  $\{N_\ell\}$ , such that,  $X_{N_\ell} \rightarrow \bar{X}$ , as  $\ell \rightarrow \infty$ , u.o.c., a.s.

**Definition 1** (fluid limit and fluid model). *A fluid limit is any limit point of the sequence of functions  $\{D(Nt)/N, t \geq 0\}$ . The fluid model is the set of these limit points.*

If  $\bar{D}(t) = (\bar{D}_1(t), \dots, \bar{D}_K(t))$  is a fluid limit, then we can define

$$\bar{Q}_k(t) = \bar{Q}_k(0) + \bar{A}_k(t) + \sum_{\ell=1}^K \bar{D}_\ell(t) p_{\ell,k} - \bar{D}_k(t), \quad k = 1, \dots, K.$$

The interpretation is easy: Since  $D(Nt)/t \rightarrow \bar{D}(t)$ , along, possibly, a subsequence, then, along the same subsequence,  $Q(Nt)/N \rightarrow \bar{Q}(t)$ . This follows from the FLLN for the arrival process and for the switching process.

**Example.** For the single-server queue, the fluid model is a collection of fluid limits indexed, say by  $c_1$  and  $c_2$ .

**Definition 2** (stability of fluid model). *We say that the fluid model is stable, if there exists a deterministic  $t_0 > 0$ , such that, for all fluid limits,  $\bar{Q}(t) = 0$  for  $t \geq t_0$ , a.s.*

To formulate a theorem, we consider the state process at the arrival epochs. So we let<sup>7</sup>  $X_n := X(T_n)$ . Then the last state component (the remaining arrival time) becomes redundant and will be omitted. Thus,  $X_n = (X_n^1, \dots, X_n^J)$ , with  $X_n^j = (Q_n^j, \zeta_n^j, \chi_n^j)$ . Define the function

$$V : ((q^j, \zeta^j, \chi^j), j = 1, \dots, J) \mapsto \sum_{j=1}^J (q^j + \chi^j).$$

**Theorem 1.** (*J. Dai, 1996*) *If the fluid model is stable, then there exists  $N_0$  such that the set  $B_{N_0} := \{x : V(x) \leq N_0\}$  is positive recurrent for  $\{X_n\}$ .*

**Remarks:**

(i) There is a number of papers where the instability conditions are analysed via fluid limits. One of the most recent is [9] where the large deviations and the martingale techniques are used.

(ii) The definition of stability of a fluid model is quite a strong one. Nevertheless, if it

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<sup>7</sup>We tacitly follow this notational convention: replacing some  $Y(t)$  by  $Y_n$  refers to sampling at time  $t = T_n$ .

holds – and it does in many important examples – then the original multiclass network is stable.

(iii) It is easy to see that the fluid model is stable in the sense of Definition 2 if and only if there exist a deterministic time  $t_0 > 0$  and a number  $\varepsilon \in (0, 1)$  such that, for all fluid limits,  $\overline{Q}(t_0) \leq 1 - \varepsilon$ , a.s.

(iv) Do we need the i.i.d. assumptions to develop the fluid approximation techniques? The answer is NO. These assumptions are needed for Theorem 1 to hold!

(v) If all fluid limits are deterministic (non-random) – like in the examples below – then the conditions for stability of the fluid model either coincide with or are close to the conditions for positive recurrence of the underlying Markov chain  $\{X_n\}$ . However, if the fluid limits remain random, stability in the sense of Definition 2 is too restrictive, and the following weaker notion of stability may be of use:

**Definition 3** (weaker notion of stability of fluid model). *The fluid model is (weakly) stable if there exist  $t_0 > 0$  and  $\varepsilon \in (0, 1)$  such that, for all fluid limits,  $\mathbf{E}\overline{Q}(t_0) \leq 1 - \varepsilon$ .*

There exist examples of stable stochastic networks whose fluid limits are a.s. not stable in the sense of Definition 2, but stable in the sense of Definition 3 (“weakly stable”) – see, e.g., [7]. The statement of Theorem 1 stays valid if one replaces the word “stable” by “weakly stable”.

*Proof of Theorem 1.* Let

$$g(x) := 2\lambda t_0 V(x), \quad h(x) := \frac{1}{2}V(x) - C\mathbf{1}(V(x) \leq N_0),$$

where  $V$  is as defined above, and  $C, N_0$  are positive constants that will be chosen suitably later. It is clear that (L0)–(L4) hold. It remains to show that the drift criterion holds. Let  $\overline{Q}$  be a fluid limit. Thus,  $Q_k(Nt)/N \rightarrow \overline{Q}_k(t)$ , along a subsequence. Hence, along the same subsequence,  $Q_{k,[N\lambda t]}/N = Q_k(T_{[N\lambda t]})/N \rightarrow \overline{Q}_k(t)$ . All limits will be taken along the subsequence referred to above and this shall not be denoted explicitly from now on. We assume that  $\overline{Q}(t) = 0$  for  $t \geq t_0$ . So,

$$\overline{\lim}_{N \rightarrow \infty} \frac{1}{N} \sum_k Q_{k,[2\lambda t_0 N]} \leq 1/2, \quad \text{a.s.} \quad (1)$$

Also,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_j \chi_n^j = 0, \quad \text{a.s.} \quad (2)$$

To see the latter, observe that, for all  $j$ ,

$$\frac{\chi_n^j}{n} \leq \frac{1}{n} \max_{k \in c(j)} \max_{1 \leq i \leq D_{k,n}+1} \sigma_k(i) \leq \sum_{k \in c(j)} \frac{D_{k,n} + 1}{n} \frac{\max_{1 \leq i \leq D_{k,n}+1} \sigma_k(i)}{D_{k,n} + 1}. \quad (3)$$

Note that

$$\frac{1}{m} \max_{1 \leq i \leq m} \sigma_k(i) \rightarrow 0, \quad \text{as } m \rightarrow \infty, \quad \text{a.s.},$$

and so

$$R_k := \sup_m \frac{1}{m} \max_{1 \leq i \leq m} \sigma_k(i) < \infty, \quad \text{a.s.}$$

The assumption that the arrival rate is finite, implies that

$$\overline{\lim}_{n \rightarrow \infty} \frac{D_{k,n} + 1}{n} < \infty, \text{ a.s.} \quad (4)$$

In case the latter quantity is positive, we have that the last fraction of (3) tends to zero. In case the latter quantity is zero then  $\chi^j(n)/n \rightarrow 0$ , because  $R_k$  is a.s. finite. We next claim that the families  $\{Q_{k,[2\lambda t_0 N]}/N\}$ ,  $\{\chi_{[2\lambda t_0 N]}^j/N\}$  are uniformly integrable. Indeed, the first one is uniformly bounded by a constant:

$$\frac{1}{N} Q_{k,[2\lambda t_0 N]} \leq \frac{1}{N} (Q_{k,0} + A(T_{[2\lambda t_0 N]})) \leq 1 + [2\lambda t_0 N]/N \leq 1 + 4\lambda t_0,$$

To see that the second family is uniformly integrable, observe that, as in (3), and if we further loosen the inequality by replacing the maximum by a sum,

$$\frac{1}{N} \chi_{[2\lambda t_0 N]}^j \leq \sum_{k \in c(j)} \frac{1}{N} \sum_{i=1}^{D_{k,[2\lambda t_0 N]}+1} \sigma_k(i),$$

where the right-hand-side can be seen to be uniformly integrable by an argument similar to the one above. From (1) and (2) and the uniform integrability we have

$$\overline{\lim}_{n \rightarrow \infty} \frac{1}{N} \left( \sum_k \mathbf{E} Q_{k,[2\lambda t_0 N]} + \sum_j \mathbf{E} \chi_{[2\lambda t_0 N]}^j \right) \leq 1/2,$$

and so there is  $N_0$ , such that, for all  $N > N_0$ ,

$$\mathbf{E} \left( \sum_k Q_{k,[2\lambda t_0 N]} + \sum_j \chi_{[2\lambda t_0 N]}^j - N \right) \leq -N/2,$$

which, using the functions introduced earlier, and the usual Markovian notation, is written as

$$\mathbf{E}_x[V(X_{g(x)}) - V(X_0)] \leq -\frac{1}{2}V(x), \quad \text{if } V(x) > N_0.$$

where the subscript  $x$  denotes the starting state, for which we had set  $N = V(x)$ . In addition,

$$\mathbf{E}_x[V(X_{g(x)}) - V(X_0)] \leq C, \quad \text{if } V(x) \leq N_0,$$

for some constant  $C < \infty$ . Thus, with  $h(x) = V(x)/2 - C\mathbf{1}(V(x) \leq N_0)$ , the last two displays combine into

$$\mathbf{E}_x[V(X_{g(x)}) - V(X_0)] \leq -h(x).$$

□

In the sequel, we present two special, but important cases, where this assumption can be verified, under usual stability conditions.

### 1.3 Multiclass queue

In this system, a special case of a multiclass queueing network, there is only one station, and  $K$  classes of customers. There is a single arrival stream  $A$  with rate  $\lambda$ . Upon arrival, a customer becomes of class  $k$  with probability  $p_k$ . Let  $A_k$  be the arrival process of class- $k$  customers. Class  $k$  customers have mean service time  $\mu_k^{-1}$ . Let  $Q_k(t)$  be the number of customers of class  $k$  in the system at time  $t$ , and let  $\chi(t)$  be the remaining service time (and hence time till departure because service discipline is non-preemptive) of the customer in service at time  $t$ . We scale according to  $N = \sum_k Q_k(0) + \chi(0)$ . We do not consider the initial time till the next arrival, because we will apply the embedding method of the previous section. The traffic intensity is  $\rho := \sum_k \lambda_k / \mu_k = \lambda \sum_k p_k / \mu_k$ . Take any subsequence such that

$$\begin{aligned} Q_k(0)/N &\rightarrow \bar{Q}_k(0), & \chi(0)/N &\rightarrow \bar{\chi}(0), \text{ a.s.}, \\ A_k(Nt)/N &\rightarrow \bar{A}_k(t) = \lambda_k t, & D_k(Nt)/N &\rightarrow \bar{D}_k(t), \text{ u.o.c., a.s.} \end{aligned}$$

That the first holds is a consequence of a FLLN. That the second holds is a consequence of Helly's extraction principle. Then  $Q(Nt)/N \rightarrow \bar{Q}(t)$ , u.o.c., a.s., and so any fluid limit satisfies

$$\begin{aligned} \bar{Q}_k(t) &= \bar{Q}_k(0) + \bar{A}_k(t) - \bar{D}_k(t), \quad k = 1, \dots, K \\ \sum_k \bar{Q}_k(0) + \bar{\chi}(0) &= 1. \end{aligned}$$

In addition, we have the following structural property for any fluid limit: define

$$\bar{I}(t) := t - \sum_k \mu_k^{-1} \bar{D}_k(t), \quad \bar{W}_k(t) := \mu_k^{-1} \bar{Q}_k(t)$$

Then  $\bar{I}$  is an increasing function, such that

$$\int_0^\infty \sum_k \bar{W}_k(t) d\bar{I}(t) = 0.$$

Hence, for any  $t$  at which the derivative exists, and at which  $\sum_k \bar{W}_k(t) > 0$ ,

$$\frac{d}{dt} \sum_k \bar{W}_k(t) = \frac{d}{dt} \left( \sum_k \mu_k^{-1} (\bar{Q}_k(0) + \bar{A}_k(t)) - t \right) - \frac{d}{dt} \bar{I}(t) = -(1 - \rho).$$

Hence, if the stability condition  $\rho < 1$  holds, then the above is strictly bounded below zero, and so, an easy argument shows that there is  $t_0 > 0$ , so that  $\sum_k \bar{W}_k(t) = 0$ , for all  $t \geq t_0$ .

N.B. This  $t_0$  is given by the formula  $t_0 = C/(1 - \rho)$  where  $C = \max\{\sum_k \mu_k^{-1} q_k + \chi : q_k \geq 0, k = 1, \dots, K, \chi \geq 0, \sum_k q_k + \chi = 1\}$ . Thus, the fluid model is stable, Theorem 1 applies, and so we have positive recurrence.

### 1.4 Jackson-type network

Here we consider another special case, where there is a customer class per station. Traditionally, when service times are exponential, we are dealing with a classical Jackson network.

This justifies our terminology “Jackson-type”, albeit, in the literature, the term “generalised Jackson” is also encountered.

Let  $\mathcal{J} := \{1, \dots, J\}$  be the set of stations (= set of classes). There is a single arrival stream  $A(t) = \mathbf{1}(\psi(0) \leq t) + \sum_{n \geq 1} \mathbf{1}(\psi(0) + T_n \leq t)$ ,  $t \geq 0$ , where  $T_n = \xi_1 + \dots + \xi_n$ ,  $n \in \mathbb{N}$ , and the  $\{\xi_n\}$  are i.i.d. positive r.v.’s with  $E\xi_1 = \lambda^{-1} \in (0, \infty)$ . Upon arrival, a customer is routed to station  $j$  with probability  $p_{0,j}$ , where  $\sum_{j=1}^J p_{0,j} = 1$ . To each station  $j$  there corresponds a random variable  $\sigma_j$  with mean  $\mu_j$ , i.i.d. copies of which are handed out as service times of customers in this station. We assume that the service discipline is non-idling, work-conserving, and non-preemptive.  $\{X(t) = [(Q^j(t), \zeta^j(t), \chi^j(t), j \in \mathcal{J}); \psi(t)], t \geq 0\}$ , as above.

The internal routing probabilities are denoted by  $p_{j,i}$ ,  $j, i \in \mathcal{J}$ : upon completion of service at station  $j$ , a customer is routed to station  $i$  with probability  $p_{j,i}$  or exits the network with probability  $1 - \sum_{i=1}^J p_{j,i}$ . We describe the (traditional) stability conditions in terms of an auxiliary Markov chain which we call  $\{Y_n\}$  and which takes values in  $\{0, 1, \dots, J, J+1\}$ , it has transition probabilities  $p_{j,i}$ ,  $j \in \{0, 1, \dots, J\}$ ,  $i \in \{1, \dots, J\}$ , and  $p_{j,J+1} = 1 - \sum_{i=1}^J p_{j,i}$ ,  $j \in \{1, \dots, J\}$ ,  $p_{J+1,J+1} = 1$ , i.e.  $J+1$  is an absorbing state. We start with  $Y_0 = 0$  and denote by  $\pi(j)$  the mean number of visits to state  $j \in \mathcal{J}$ :

$$\pi(j) = E \sum_n \mathbf{1}(Y_n = j) = \sum_n P(Y_n = j).$$

Firstly we assume (and this is no loss of generality) that  $\pi(j) > 0$  for all  $j \in \mathcal{J}$ . Secondly, we assume that

$$\max_{j \in \mathcal{J}} \pi(j) \mu_j^{-1} < \lambda^{-1}.$$

Now scale according to  $N = \sum_{j=1}^J [Q_j(0) + \chi_j(0)]$ . Again, due to our embedding technique, we assume at the outset that  $\psi(0) = 0$ . By applying the FLLN it is seen that any fluid limit satisfies

$$\begin{aligned} \bar{Q}_j(t) &= \bar{Q}_j(0) + \bar{A}_j(t) + \sum_{i=1}^J \bar{D}_i(t) p_{i,j} - \bar{D}_j(t), \quad j \in \mathcal{J} \\ \sum_j [\bar{Q}_j(0) + \bar{\chi}_j(0)] &= 1, \\ \bar{A}_j(t) &= \lambda_j t = \lambda p_{0,j} t, \quad \bar{D}_j(t) = \mu_j(t - \bar{I}_j(t)), \end{aligned}$$

where  $\bar{I}_j$  is an increasing function, representing cumulative idleness at station  $j$ , such that

$$\sum_{j=1}^J \int_0^\infty \bar{Q}_j(t) d\bar{I}_j(t) = 0.$$

We next show that the fluid model is stable, i.e., that there exists a  $t_0 > 0$  such that  $\bar{Q}(t) = 0$  for all  $t \geq t_0$ .

We base this on the following facts: If a function  $g : \mathbb{R} \rightarrow \mathbb{R}^n$  is Lipschitz then it is a.e. differentiable. A point of differentiability of  $g$  (in the sense that the derivative of all its coordinates exists) will be called “regular”. Suppose then that  $g$  is Lipschitz with  $\sum_{i=1}^n g_i(0) =: |g(0)| > 0$  and  $\varepsilon > 0$  such that ( $t$  regular and  $|g(t)| > 0$ ) imply  $|g(t)' \leq -\varepsilon$ ;

then  $|g(t)| = 0$  for all  $t \geq |g(0)|/\varepsilon$ . Finally, if  $h : \mathbb{R} \rightarrow \mathbb{R}$  is a non-negative Lipschitz function and  $t$  a regular point at which  $h(t) = 0$  then necessarily  $h'(t) = 0$  (see the Appendix for details).

We apply these to the Lipschitz function  $\bar{Q}$ . It is sufficient to show that for any  $\mathcal{I} \subseteq \mathcal{J}$  there exists  $\varepsilon = \varepsilon(\mathcal{I}) > 0$  such that, for any regular  $t$  with  $\min_{i \in \mathcal{I}} \bar{Q}_i(t) > 0$  and  $\max_{i \in \mathcal{J} - \mathcal{I}} \bar{Q}_i(t) = 0$ , we have  $|\bar{Q}(t)'| \leq -\varepsilon$ . Suppose first that  $\mathcal{I} = \mathcal{J}$ . That is, suppose  $\bar{Q}_j(t) > 0$  for all  $j \in \mathcal{J}$ , and  $t$  a regular point. Then  $\bar{Q}_j(t)' = \lambda_j + \sum_{i=1}^J \mu_i p_{i,j} - \mu_j$  and so  $|\bar{Q}_j(t)'| = \lambda - \sum_{j=1}^J \sum_{i=1}^J \mu_i p_{i,j} - \sum_{j=1}^J \mu_j = \lambda - \sum_{i=1}^J \mu_i p_{i,J+1} =: -\varepsilon(\mathcal{J})$ . But  $\mu_i > \pi(i)\lambda$  and so  $\varepsilon(\mathcal{J}) > \lambda(1 - \sum_{i=1}^J \pi(i)p_{i,J+1}) = 0$ , where the last equality follows from  $\sum_{i=1}^J \pi(i)p_{i,J+1} = \sum_{i=1}^J \sum_n \mathbf{P}(Y_n = i, Y_{n+1} = J+1) = \sum_n \mathbf{P}(Y_n \neq J+1, Y_{n+1} = J+1) = 1$ .

Next consider  $\mathcal{I} \subset \mathcal{J}$ . Consider an auxiliary Jackson-type network that is derived from the original one by  $\sigma_j = 0$  for all  $j \in \mathcal{J} - \mathcal{I}$ . It is then clear that this network has routing probabilities  $p_{i,j}^{\mathcal{I}}$  that correspond to the Markov chain  $\{Y_n^{\mathcal{I}}\}$  being a subsequence of  $\{Y_n\}$  at those epochs  $n$  for which  $Y_n \in \mathcal{I} \cup \{J+1\}$ . Let  $\pi^{\mathcal{I}}(i)$  the mean number of visits to state  $i \in \mathcal{I}$  by this embedded chain. Clearly,  $\pi^{\mathcal{I}}(i) = \pi(i)$ , for all  $i \in \mathcal{I}$ . So the stability condition  $\max_{i \in \mathcal{I}} \pi(i)\mu_i < \lambda^{-1}$  is a trivial consequence of the stability condition for the original network. Also, the fluid model for the auxiliary network is easily derived from that of the original one. Assume then  $t$  is a regular point with  $\min_{i \in \mathcal{I}} \bar{Q}_i(t) > 0$  and  $\max_{i \in \mathcal{J} - \mathcal{I}} \bar{Q}_i(t) = 0$ . Then  $|\bar{Q}_j(t)'| = 0$  for all  $j \in \mathcal{J} - \mathcal{I}$ . By interpreting this as a statement about the fluid model of the auxiliary network, in other words that all queues of the fluid model of the auxiliary network are positive at time  $t$ , we have, precisely as in the previous paragraph, that  $\bar{Q}_j(t)' = \lambda p_{0,j}^{\mathcal{I}} + \sum_{i \in \mathcal{I}} \mu_i p_{i,j}^{\mathcal{I}} - \mu_j$ , for all  $j \in \mathcal{I}$ , and so  $|\bar{Q}(t)'| = \lambda - \sum_{i \in \mathcal{I}} \mu_i p_{i,J+1}^{\mathcal{I}} =: -\varepsilon(\mathcal{I})$ . As before,  $\varepsilon(\mathcal{I}) > \lambda(1 - \sum_{i \in \mathcal{I}} \pi(i)p_{i,J+1}^{\mathcal{I}}) = 0$ .

We have thus proved that, with  $\varepsilon := \min_{\mathcal{I} \subseteq \mathcal{J}} \varepsilon(\mathcal{I})$ , for any regular point  $t$ , if  $|\bar{Q}(t)'| > 0$ , then  $|\bar{Q}(t)| \leq -\varepsilon$ . Hence the fluid model is stable.

We considered multiclass networks with single-server stations.

**Exercise 1.** Consider a two-server FCFS queue with i.i.d. inter-arrival and i.i.d. service times queue, and introduce a fluid model for it. Then find stability conditions.

**Exercise 2.** More generally, study a multi-server queue.

**Exercise 3.** Find stability conditions for a tandem of two 2-server queues.

**Exercise 4.** Study a tandem of two 2-server queues with feedback: upon service completion at station 2, a customer returns to station 1 with probability  $p$  and leaves the network otherwise.

## 2 Inducing (second) vector field

In this section, we consider only a particular class of models: Markov chains in the positive quadrant  $\mathcal{ZR}^2$ . An analysis of more general models may be found, e.g., in [1, 6, 12]. We follow here [1], Chapter 7.

Let  $\{X_n\}$  be a Markov chain in  $\mathcal{ZR}^2$  with initial state  $X_0$ . For  $(x, y) \in \mathcal{R}^2$ , let a random

vector  $\xi_{x,y}$  have a distribution

$$\mathbf{P}(\xi_{x,y} \in \cdot) = \mathbf{P}(X_1 - X_0 \in \cdot \mid X_0 = (x, y))$$

and let

$$a_{x,y} = \mathbf{E}\xi_{x,y} \equiv (a_{x,y}^{(1)}, a_{x,y}^{(2)})$$

be a 1-step mean drift vector from point  $(x, y)$ .

Assume that random variables  $\{\xi_{x,y}\}$  are uniformly integrable and that a Markov chain is *asymptotically homogeneous* in the following sense: first,

$$\xi_{x,y} \rightarrow \xi \quad \text{weakly as } x, y \rightarrow \infty,$$

then  $a_{x,y} \rightarrow a = (a^{(1)}, a^{(2)}) = \mathbf{E}\xi$ . Also,

$$\xi_{x,y} \rightarrow \xi_{x,\infty} \quad \text{weakly as } y \rightarrow \infty, \quad \forall x,$$

then  $a_{x,y} \rightarrow a_{x,\infty} = \mathbf{E}\xi_{x,\infty}$ ; and

$$\xi_{x,y} \rightarrow \xi_{\infty,y} \quad \text{weakly as } x \rightarrow \infty, \quad \forall y,$$

then  $a_{x,y} \rightarrow a_{\infty,y} = \mathbf{E}\xi_{\infty,y}$ . Note also that  $a_{x,\infty} \rightarrow a$  as  $x \rightarrow \infty$  and  $a_{\infty,y} \rightarrow a$  as  $y \rightarrow \infty$ .

Consider a homogeneous Markov chain  $V_n^{(1)}$  on  $\mathbb{R}$  with distributions of increments

$$\mathbf{P}_v(V_1^{(1)} - V_0^{(1)} \in \cdot) = \mathbf{P}(\xi_{\infty,v}^{(1)} \in \cdot)$$

and a homogeneous Markov chain  $V_n^{(2)}$  on  $\mathbb{R}$  with distributions of increments

$$\mathbf{P}_y(V_1^{(2)} - V_0^{(2)} \in \cdot) = \mathbf{P}(\xi_{\infty,v}^{(2)} \in \cdot)$$

We also need an extra

**Assumption.** For  $i = 1, 2$ , if  $a^{(i)} < 0$ , then a Markov chain  $\{V_n^{(i)}\}$  converges to a stationary distribution  $\pi^{(i)}$ . In this case, let

$$c^{(i)} = \int_0^\infty \pi^{(i)}(dv) a^{(3-i)}(\dots)$$

Here (...) means  $(v, \infty)$  if  $i = 1$  and  $(\infty, v)$  if  $i = 2$ .

**Theorem 2.** Assume that  $a^{(1)} \neq 0$  and  $a^{(2)} \neq 0$ . Assume further that  $\min(a^{(1)}, a^{(2)}) < 0$  and, for  $i = 1, 2$ , if  $a^{(i)} < 0$ , then  $c^{(i)} < 0$ . Then a Markov chain  $X_n$  is positive recurrent.

PROOF is omitted. We provide verbally some intuition instead.

**Example.** Consider a tandem of two queues with state-dependent feedback. Assume that all driving random variables are mutually independent and have exponential distributions: – an exogenous input is a Poisson process with parameter  $\lambda$ , this means that the interarrival times are i.i.d.  $\text{Exp}(\lambda)$ ; – service time at station  $i = 1, 2$  have exponential distribution with parameter  $\mu_i$ .

In addition, after a service completion at station 2, a customer returns to station 1 with probability  $p_{n_1, n_2}$  and leaves the network otherwise. Here  $n_i$  is a number of customers at station  $i$  prior to completion of service..

After doing embedding (or uniformisation), we get a discrete time Markov chains. For this Markov chain, one of three events may happen: either a new customer arrives to station 1 (with prob  $\lambda/(\lambda + \mu_1 + \mu_2)$ ) or a service is completed at station 1 ( w.p.  $\mu_1/(\lambda + \mu_1 + \mu_2)$ , this will be an artificial service if station 1 is empty) or a service is completed at station 2 (again it may be an artificial service, and if not, then a customer returns to station 1 with probability  $p(\cdot, \cdot)$ ). Thus, only moves to some neighbouring states are possible. Given that a Markov chain is at state  $(i, j)$ ,

(a) if  $i > 0, j > 0$ , then

$$\begin{aligned} P((i, j), (i + 1, j)) &= \frac{\lambda}{\lambda + \mu_1 + \mu_2}, & P((i, j), (i - 1, j + 1)) &= \frac{\mu_1}{\lambda + \mu_1 + \mu_2}, \\ P((i, j), (i + 1, j - 1)) &= \frac{\mu_2 p(i, j)}{\lambda + \mu_1 + \mu_2}, & P((i, j), (i, j - 1)) &= \frac{\mu_2(1 - p(i, j))}{\lambda + \mu_1 + \mu_2} \end{aligned}$$

(b) if  $i > j = 0$ , then

$$\begin{aligned} P((i, 0), (i + 1, 0)) &= \frac{\lambda}{\lambda + \mu_1 + \mu_2}, & P((i, 0), (i - 1, 1)) &= \frac{\mu_1}{\lambda + \mu_1 + \mu_2}, \\ P((i, 0), (i, 0)) &= \frac{\mu_2}{\lambda + \mu_1 + \mu_2}, \end{aligned}$$

(c) if  $j > i = 0$ , then

$$\begin{aligned} P((0, j), (1, j)) &= \frac{\lambda}{\lambda + \mu_1 + \mu_2}, & P((0, j), (0, j)) &= \frac{\mu_1}{\lambda + \mu_1 + \mu_2}, \\ P((0, j), (1, j - 1)) &= \frac{\mu_2 p(i, j)}{\lambda + \mu_1 + \mu_2}, & P((0, j), (0, j - 1)) &= \frac{\mu_2(1 - p(i, j))}{\lambda + \mu_1 + \mu_2} \end{aligned}$$

(d) finally, if  $i = j = 0$ , then

$$P((0, 0), (1, 0)) = 1 - P((0, 0), (0, 0)) = \frac{\lambda}{\lambda + \mu_1 + \mu_2}.$$

This is *asymptotically* and, moreover, *partially homogeneous* Markov chain.

We may consider first a particular case when the probabilities  $p(n_1, n_2)$  depend on  $n_2$  only. Assume that there exists a limit  $p = \lim_{n_2 \rightarrow \infty} p(n_2)$ .

The rest of the analysis of this example is a “small research project” to you:

**Exercise 5.** Find stability conditions in terms of  $\lambda$ ,  $\mu_1$ , and  $\mu_2$ .

**Exercise 6.** Consider then the case where the probabilities  $p(n_1, n_2)$  depend on  $n_1$  only. Assuming the existence of the limit  $p = \lim_{n_2 \rightarrow \infty} p(n_2)$ , find stability conditions.

**Exercise 7.** Consider finally a general case where the probabilities  $p(n_1, n_2)$  may depend on both  $n_1$  and  $n_2$ .

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